

Modeling and Verification of Probabilistic Systems

Lecture 14: Continuous-Time Markov Chains

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June 20, 2011

Overview

1 Negative exponential distribution

2 Continuous-time Markov chains

3 Transient distribution

4 Summary

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Time in discrete-time Markov chains

The advance of time in DTMCs

- ▶ Time in a DTMC proceeds in **discrete steps**
- ▶ Two possible interpretations:
 1. accurate model of (discrete) time units
 - ▶ e.g., clock ticks in model of an embedded device
 2. time-abstract
 - ▶ no information assumed about the time transitions take
- ▶ State residence time is **geometrically** distributed

Continuous-time Markov chains

- ▶ dense model of time
- ▶ transitions can occur at any (real-valued) time instant
- ▶ state residence time is **(negative) exponentially** distributed

Continuous random variables

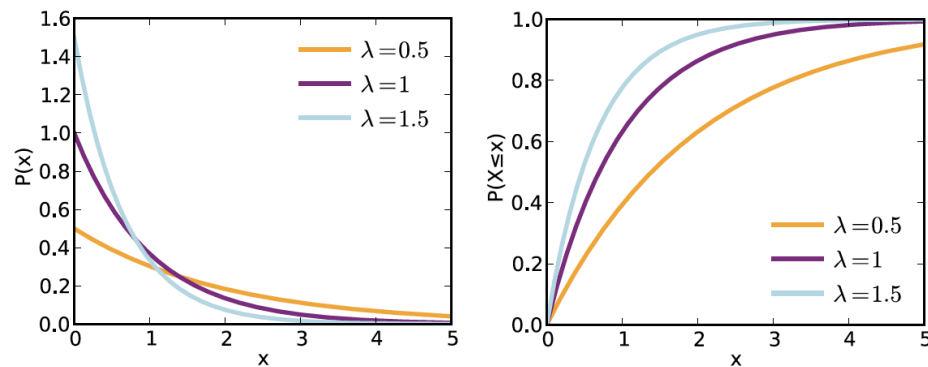
- ▶ X is a random variable (r.v., for short)
 - ▶ on a sample space with probability measure Pr
 - ▶ assume the set of possible values that X may take is dense
- ▶ X is *continuously distributed* if there exists a function $f(x)$ such that:

$$F_X(d) = Pr\{X \leq d\} = \int_{-\infty}^d f(x) dx \quad \text{for each real number } d$$

where f satisfies: $f(x) \geq 0$ for all x and $\int_{-\infty}^{\infty} f(x) dx = 1$

- ▶ $F_X(d)$ is the *(cumulative) probability distribution function*
- ▶ $f(x)$ is the *probability density function*

Exponential pdf and cdf



The higher λ , the faster the cdf approaches 1.

Negative exponential distribution

Density of exponential distribution

The density of an *exponentially distributed* r.v. Y with *rate* $\lambda \in \mathbb{R}_{>0}$ is:

$$f_Y(x) = \lambda \cdot e^{-\lambda \cdot x} \quad \text{for } x > 0 \quad \text{and } f_Y(x) = 0 \text{ otherwise}$$

The cumulative distribution of r.v. Y with rate $\lambda \in \mathbb{R}_{>0}$ is:

$$F_Y(d) = \int_0^d \lambda \cdot e^{-\lambda \cdot x} dx = [-e^{-\lambda \cdot x}]_0^d = 1 - e^{-\lambda \cdot d}.$$

The rate $\lambda \in \mathbb{R}_{>0}$ uniquely determines an exponential distribution.

Variance and expectation

Let r.v. Y be exponentially distributed with rate $\lambda \in \mathbb{R}_{>0}$. Then:

- ▶ Expectation $E[Y] = \int_0^{\infty} x \cdot \lambda \cdot e^{-\lambda \cdot x} dx = \frac{1}{\lambda}$
- ▶ Variance $Var[Y] = \int_0^{\infty} (x - E[X])^2 \lambda \cdot e^{-\lambda \cdot x} dx = \frac{1}{\lambda^2}$

Why exponential distributions?

- ▶ Are *adequate* for many real-life phenomena
 - ▶ the time until a radioactive particle decays
 - ▶ the time between successive car accidents
 - ▶ inter-arrival times of jobs, telephone calls in a fixed interval
- ▶ Are the continuous counterpart of the *geometric* distribution
- ▶ Heavily used in physics, performance, and reliability analysis
- ▶ Can *approximate* general distributions arbitrarily closely
- ▶ Yield a *maximal entropy* if only the mean is known

Memoryless property

Theorem

1. For any exponentially distributed random variable X :

$$Pr\{X > t + d \mid X > t\} = Pr\{X > d\} \text{ for any } t, d \in \mathbb{R}_{\geq 0}.$$

2. Any cdf which is memoryless is a negative exponential one.

Proof:

Proof of 1. : Let λ be the rate of X 's distribution. Then we derive:

$$\begin{aligned} Pr\{X > t + d \mid X > t\} &= \frac{Pr\{X > t + d \cap X > t\}}{Pr\{X > t\}} = \frac{Pr\{X > t + d\}}{Pr\{X > t\}} \\ &= \frac{e^{-\lambda \cdot (t+d)}}{e^{-\lambda \cdot t}} = e^{-\lambda \cdot d} = Pr\{X > d\}. \end{aligned}$$

Proof of 2. : By contraposition, using the total law of probability.

Proof

Let λ (μ) be the rate of X 's (Y 's) distribution. Then we derive:

$$\begin{aligned} Pr\{\min(X, Y) \leq t\} &= Pr_{X,Y}\{(x, y) \in \mathbb{R}_{\geq 0}^2 \mid \min(x, y) \leq t\} \\ &= \int_0^\infty \left(\int_0^\infty \mathbf{1}_{\min(x,y) \leq t}(x, y) \cdot \lambda e^{-\lambda x} \cdot \mu e^{-\mu y} dy \right) dx \\ &= \int_0^t \int_x^\infty \lambda e^{-\lambda x} \cdot \mu e^{-\mu y} dy dx + \int_0^t \int_y^\infty \lambda e^{-\lambda x} \cdot \mu e^{-\mu y} dx dy \\ &= \int_0^t \lambda e^{-\lambda x} \cdot e^{-\mu x} dx + \int_0^t e^{-\lambda y} \cdot \mu e^{-\mu y} dy \\ &= \int_0^t \lambda e^{-(\lambda+\mu)x} dx + \int_0^t \mu e^{-(\lambda+\mu)y} dy \\ &= \int_0^t (\lambda + \mu) \cdot e^{-(\lambda+\mu)z} dz = 1 - e^{-(\lambda+\mu)t} \end{aligned}$$

Closure under minimum

Minimum closure theorem

For independent, exponentially distributed random variables X and Y with rates $\lambda, \mu \in \mathbb{R}_{>0}$, the r.v. $\min(X, Y)$ is exponentially distributed with rate $\lambda + \mu$, i.e.,:

$$Pr\{\min(X, Y) \leq t\} = 1 - e^{-(\lambda+\mu)t} \text{ for all } t \in \mathbb{R}_{\geq 0}.$$

Closure under minimum

Minimum closure theorem for several exponentially distributed r.v.'s

For independent, exponentially distributed random variables X_1, X_2, \dots, X_n with rates $\lambda_1, \lambda_2, \dots, \lambda_n \in \mathbb{R}_{>0}$ the r.v. $\min(X_1, X_2, \dots, X_n)$ is exponentially distributed with rate $\sum_{0 < i \leq n} \lambda_i$, i.e.,:

$$Pr\{\min(X_1, X_2, \dots, X_n) \leq t\} = 1 - e^{-\sum_{0 < i \leq n} \lambda_i t} \text{ for all } t \in \mathbb{R}_{\geq 0}.$$

Proof:

Generalization of the proof for the case of two exponential distributions.

Winning the race with two competitors

The minimum of two exponential distributions

For independent, exponentially distributed random variables X and Y with rates $\lambda, \mu \in \mathbb{R}_{>0}$, it holds:

$$Pr\{X \leq Y\} = \frac{\lambda}{\lambda + \mu}.$$

Winning the race with many competitors

The minimum of several exponentially distributed r.v.'s

For independent, exponentially distributed random variables X_1, X_2, \dots, X_n with rates $\lambda_1, \lambda_2, \dots, \lambda_n \in \mathbb{R}_{>0}$ it holds:

$$Pr\{X_i = \min(X_1, \dots, X_n)\} = \frac{\lambda_i}{\sum_{j=1}^n \lambda_j}.$$

Proof:

Generalization of the proof for the case of two exponential distributions.

Proof

Let λ (μ) be the rate of X 's (Y 's) distribution. Then we derive:

$$\begin{aligned} Pr\{X \leq Y\} &= Pr_{X,Y}\{(x,y) \in \mathbb{R}_{\geq 0}^2 \mid x \leq y\} \\ &= \int_0^\infty \mu e^{-\mu y} \left(\int_0^y \lambda e^{-\lambda x} dx \right) dy \\ &= \int_0^\infty \mu e^{-\mu y} (1 - e^{-\lambda y}) dy \\ &= 1 - \int_0^\infty \mu e^{-\mu y} \cdot e^{-\lambda y} dy = 1 - \int_0^\infty \mu e^{-(\mu+\lambda)y} dy \\ &= 1 - \frac{\mu}{\mu+\lambda} \cdot \underbrace{\int_0^\infty (\mu+\lambda) e^{-(\mu+\lambda)y} dy}_{=1} \\ &= 1 - \frac{\mu}{\mu+\lambda} = \frac{\lambda}{\mu+\lambda} \end{aligned}$$

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Continuous-time Markov chain

Continuous-time Markov chain

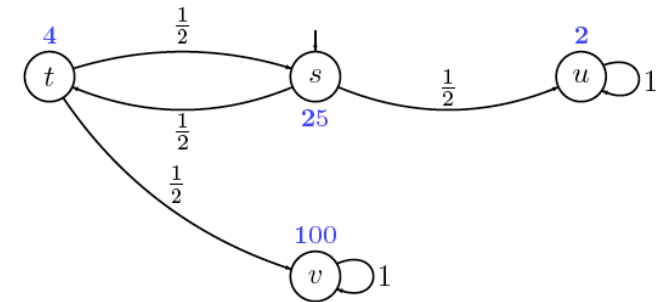
A CTMC is a tuple $(S, \mathbf{P}, r, \ell_{\text{init}}, AP, L)$ where

- ▶ $(S, \mathbf{P}, \ell_{\text{init}}, AP, L)$ is a DTMC, and
- ▶ $r : S \rightarrow \mathbb{R}_{>0}$, the exit-rate function

Interpretation

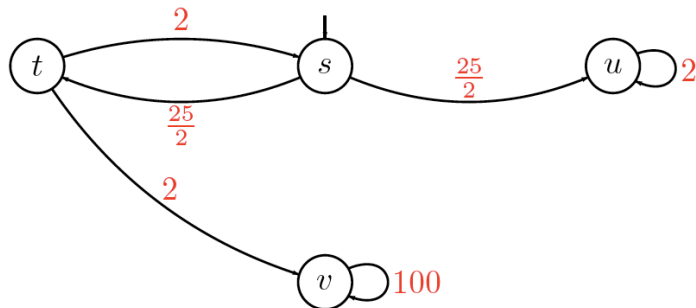
- ▶ residence time in state s is exponentially distributed with rate $r(s)$.
- ▶ phrased alternatively, the average residence time of state s is $\frac{1}{r(s)}$.
- ▶ thus, the higher the rate $r(s)$, the shorter the average residence time in s .

Example



$$r(s) = 25, r(t) = 4, r(u) = 2 \text{ and } r(v) = 100$$

Example: a classical perspective



$$r(s) = 25, r(t) = 4, r(u) = 2 \text{ and } r(v) = 100$$

$$\text{The transition rate } \mathbf{R}(s, s') = \mathbf{P}(s, s') \cdot r(s)$$

We use $(S, \mathbf{P}, r, \ell_{\text{init}}, AP, L)$ and $(S, \mathbf{R}, \ell_{\text{init}}, AP, L)$ interchangeably.

CTMC semantics by example

CTMC semantics

- ▶ Transition $s \rightarrow s' := \text{r.v. } X_{s,s'}$ with rate $\mathbf{R}(s, s')$
- ▶ Probability to go from state s_0 to, say, state s_2 is:

$$\begin{aligned} & \Pr\{X_{s_0, s_2} \leq X_{s_0, s_1} \cap X_{s_0, s_2} \leq X_{s_0, s_3}\} \\ &= \frac{\mathbf{R}(s_0, s_2)}{\mathbf{R}(s_0, s_1) + \mathbf{R}(s_0, s_2) + \mathbf{R}(s_0, s_3)} = \frac{\mathbf{R}(s_0, s_2)}{r(s_0)} \end{aligned}$$

- ▶ Probability of staying at most t time in s_0 is:

$$\begin{aligned} & \Pr\{\min(X_{s_0, s_1}, X_{s_0, s_2}, X_{s_0, s_3}) \leq t\} \\ &= 1 - e^{-(\mathbf{R}(s_0, s_1) + \mathbf{R}(s_0, s_2) + \mathbf{R}(s_0, s_3)) \cdot t} = 1 - e^{-r(s_0) \cdot t} \end{aligned}$$

CTMC semantics

Enabledness

The probability that transition $s \rightarrow s'$ is *enabled* in $[0, t]$ is $1 - e^{-R(s,s') \cdot t}$.

State-to-state timed transition probability

The probability to *move* from non-absorbing s to s' in $[0, t]$ is:

$$\frac{R(s, s')}{r(s)} \cdot (1 - e^{-r(s) \cdot t}).$$

Residence time distribution

The probability to *take some* outgoing transition from s in $[0, t]$ is:

$$\int_0^t r(s) \cdot e^{-r(s) \cdot x} dx = 1 - e^{-r(s) \cdot t}$$

CTMC semantics

Residence time distribution

The probability to *take some* outgoing transition from s in $[0, t]$ is:

$$\int_0^t r(s) \cdot e^{-r(s) \cdot x} dx = 1 - e^{-r(s) \cdot t}$$

Proof:

On the blackboard.

CTMC semantics

State-to-state timed transition probability

The probability to *move* from non-absorbing s to s' in $[0, t]$ is:

$$\frac{R(s, s')}{r(s)} \cdot (1 - e^{-r(s) \cdot t}).$$

Proof:

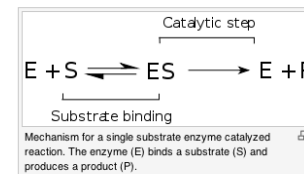
On the blackboard.

Enzyme-catalysed substrate conversion

Kinetics

[\[edit\]](#)

Main article: [Enzyme kinetics](#)



Enzyme kinetics is the investigation of how enzymes bind substrates and turn them into products. The rate data used in kinetic analyses are commonly obtained from [enzyme assays](#), where since the 90s, the dynamics of many enzymes are studied on the level of [individual molecules](#).

In 1902 Victor Henri^[57] proposed a quantitative theory of enzyme kinetics, but his experimental data were not useful because the significance of the hydrogen ion concentration was not yet appreciated. After [Peter Lauritz Sørensen](#) had defined the logarithmic pH-scale and introduced the concept of buffering in 1909^[58] the German chemist [Leonor Michaelis](#) and his Canadian postdoc [Maud Leonora Menten](#) repeated Henri's experiments and confirmed his equation which is referred to as [Henri-Michaelis-Menten kinetics](#) (termed also [Michaelis-Menten kinetics](#)).^[59]

Their work was further developed by [G. E. Briggs](#) and [J. B. S. Haldane](#), who derived kinetic

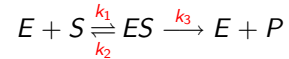
equations that are still widely considered today a starting point in solving enzymatic activity.^[60]

The major contribution of Henri was to think of enzyme reactions in two stages. In the first, the substrate binds reversibly to the enzyme, forming the enzyme-substrate complex. This is sometimes called the Michaelis complex. The enzyme then catalyzes the chemical step in the reaction and releases the product. Note that the simple [Michaelis-Menten mechanism](#) for the enzymatic activity is considered today a basic idea, where many examples show that the enzymatic activity involves structural dynamics. This is incorporated in the enzymatic mechanism while introducing several Michaelis-Menten pathways that are connected with fluctuating rates ^{[44][45][46]}. Nevertheless, there is a mathematical relation connecting the behavior obtained from the basic Michaelis-Menten mechanism (that was indeed proved correct in many experiments) with the generalized Michaelis-Menten mechanisms involving dynamics and activity; ^[61] this means that the measured activity of enzymes on the level of many enzymes may be explained with the simple Michaelis-Menten equation, yet, the actual activity of enzymes is richer and involves structural dynamics.

Source: wikipedia (June 2011)

Stochastic chemical kinetics

- Types of reaction described by **stoichiometric equations**:



- N different types of molecules that **randomly collide**
where state $X(t) = (x_1, \dots, x_N)$ with $x_i = \#$ molecules of sort i

- Reaction probability** within infinitesimal interval $[t, t+\Delta)$:

$$\alpha_m(\vec{x}) \cdot \Delta = Pr\{\text{reaction } m \text{ in } [t, t+\Delta) \mid X(t) = \vec{x}\} \text{ where}$$

$$\alpha_m(\vec{x}) = k_m \cdot \# \text{ possible combinations of reactant molecules in } \vec{x}$$

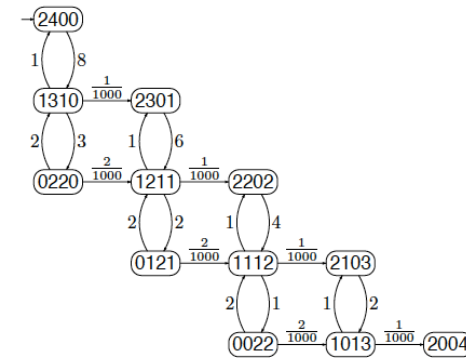
- This process is a **continuous-time Markov chain**.

CTMCs are omnipresent!

- Markovian queueing networks (Kleinrock 1975)
- Stochastic Petri nets (Molloy 1977)
- Stochastic activity networks (Meyer & Sanders 1985)
- Stochastic process algebra (Herzog et al., Hillston 1993)
- Probabilistic input/output automata (Smolka et al. 1994)
- Calculi for biological systems (Priami et al., Cardelli 2002)

CTMCs are one of the most prominent models in performance analysis

Enzyme-catalyzed substrate conversion as a CTMC



| States: | init | goal |
|------------|------|------|
| enzymes | 2 | 2 |
| substrates | 4 | 0 |
| complex | 0 | 0 |
| products | 0 | 4 |

Transitions: $E + S \xrightleftharpoons[1]{1} C \xrightarrow{0.001} E + P$

e.g., $(x_E, x_S, x_C, x_P) \xrightarrow{0.001 \cdot x_C} (x_E + 1, x_S, x_C - 1, x_P + 1)$ for $x_C > 0$

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Transient distribution of a CTMC

Transient state probability

Let $X(t)$ denote the state of a CTMC at time $t \in \mathbb{R}_{\geq 0}$. The probability to be in state s at time t is defined by:

$$\begin{aligned} p_s(t) &= \Pr\{X(t) = s\} \\ &= \sum_{s' \in S} \Pr\{X(0) = s'\} \cdot \Pr\{X(t) = s \mid X(0) = s'\} \end{aligned}$$

Theorem: transient distribution as linear differential equation

The **transient** probability vector $\underline{p}(t) = (p_{s_1}(t), \dots, p_{s_k}(t))$ satisfies:

$$\underline{p}'(t) = \underline{p}(t) \cdot (\mathbf{R} - \mathbf{r}) \quad \text{given} \quad \underline{p}(0)$$

where \mathbf{r} is the diagonal matrix of vector \underline{r} .

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Transient distribution theorem

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Proof:

On the blackboard.

Summary

Main points

- ▶ Exponential distributions are closed under minimum.
- ▶ The probability to win a race amongst several exponential distributions only depends on their rates.
- ▶ A CTMC is a DTMC where state residence times are exponentially distributed.
- ▶ CTMC semantics distinguishes between enabledness and taking a transition.
- ▶ Transient distribution are obtained by solving a system of linear differential equations.
- ▶ CTMCs are frequently used as semantical model for high-level formalisms.